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Sensitivity Equations Provide More Robust Gradients and Faster Computation of the FOCE Approximation to the Population Likelihood

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Background

The first order conditional estimation (FOCE) method [1] is still one of the parameter estimation workhorses for nonlinear mixed effects (NLME) modeling used in population pharmacokinetics and pharmacodynamics [2]. We propose an novel implementation of the FOCE and FOCEI methods where instead of obtaining the gradients needed for the two levels of quasi-Newton optimizations from the standard finite difference approximation, gradients are computed using so called sensitivity equations [3].

The Approximate Population Likelihood

The state-space model for a single individual is described by a system of ordinary differential equations and a corresponding set of measurement equations

The Outer Optimization Problem

The outer optimization problem consists of finding the θ that maximizes the log-likelihood. The m^{th} component of the gradient of the log-likelihood wrt θ

$$\frac{d\mathbf{x}_{i}(t)}{dt} = \mathbf{f}(\mathbf{x}_{i}(t), t, \mathbf{Z}_{i}(t), \boldsymbol{\theta}, \boldsymbol{\eta}_{i}) \qquad \mathbf{y}_{ij} = \mathbf{h}(\mathbf{x}_{ij}, t_{j_{i}}, \mathbf{Z}_{i}(t_{j_{i}}), \boldsymbol{\theta}, \boldsymbol{\eta}_{i}) + \mathbf{e}_{ij}$$
$$\mathbf{y}_{ij} \in \mathbf{N}(\mathbf{0}, \mathbf{R}_{ij}(\mathbf{x}_{ij}, t_{j_{i}}, \mathbf{Z}_{i}(t_{j_{i}}), \boldsymbol{\theta}, \boldsymbol{\eta}_{i}))$$
$$\mathbf{x}_{i}(t_{0}) = \mathbf{x}_{0i}(\mathbf{Z}_{i}(t_{0}), \boldsymbol{\theta}, \boldsymbol{\eta}_{i}) \qquad \hat{\mathbf{y}}_{ij} = \mathbf{E}[\mathbf{y}_{ij}]$$

where indices *i* and *j* denote individuals and observations, respectively. Furthermore, θ are fixed effects parameters, $Z_i(t_j)$ are covariates, $\eta_i \sim N(0, \Omega)$ are random effect parameters, and R_{ij} are measurement error covariance matrices.

Given a set of experimental observations, d_{ij} , for the individuals i = 1, ..., N at the time points t_{ij} , where $j = 1, ..., n_i$, we define the residuals $\epsilon_{ij} = d_{ij} - \hat{y}_{ij}$

The approximate log-likelihood function is obtained using the Laplacian approximation, which involves a second order Taylor expansion wrt η_i of l_i around points η_i^* that maximizes the individual l_i .

$$\log L(\boldsymbol{\theta}) \approx \log L_F(\boldsymbol{\theta}) = \sum_{i=1}^N \left(l_i(\boldsymbol{\eta}_i^*) - \frac{1}{2} \log \det \left[\frac{-\mathbf{H}_i(\boldsymbol{\eta}_i^*)}{2\pi} \right] \right)$$

where

$$\frac{d\log L_F}{d\theta_m} = \sum_{i=1}^N \left(\frac{dl_i(\boldsymbol{\eta}_i^*)}{d\theta_m} - \frac{1}{2} \operatorname{tr} \left[\mathbf{H}_i^{-1}(\boldsymbol{\eta}_i^*) \frac{d\mathbf{H}_i(\boldsymbol{\eta}_i^*)}{d\theta_m} \right] \right)$$

where the total derivatives of l_i and H_i wrt θ can be expressed in terms of solutions to the sensitivity differential equations, e.g.,

$$\begin{split} \frac{dl_{i}(\boldsymbol{\eta}_{i}^{*})}{d\theta_{m}} &= \frac{dl_{i}(\boldsymbol{\eta}_{i})}{d\theta_{m}} \Big|_{\boldsymbol{\eta}_{i} = \boldsymbol{\eta}_{i}^{*}(\boldsymbol{\theta})} = \left[-\frac{1}{2} \sum_{j=1}^{n_{i}} \left(2\epsilon_{ij}^{T} \mathbf{R}_{ij}^{-1} \frac{d\epsilon_{ij}}{d\theta_{m}} - \epsilon_{ij}^{T} \mathbf{R}_{ij}^{-1} \frac{d\mathbf{R}_{ij}}{d\theta_{m}} \mathbf{R}_{ij}^{-1} \epsilon_{ij} \right. \\ &+ \mathrm{tr} \left[\mathbf{R}_{ij}^{-1} \frac{d\mathbf{R}_{ij}}{d\theta_{m}} \right] \right) + \frac{1}{2} \boldsymbol{\eta}_{i} \, \Omega^{-1} \frac{d\Omega}{d\theta_{m}} \Omega^{-1} \boldsymbol{\eta}_{i} - \frac{1}{2} \mathrm{tr} \left[\Omega^{-1} \frac{d\Omega}{d\theta_{m}} \right] \right]_{\boldsymbol{\eta}_{i} = \boldsymbol{\eta}_{i}^{*}(\boldsymbol{\theta})} \\ &\frac{d\epsilon_{ij}^{*}}{d\theta_{m}} = \frac{d\epsilon_{ij}}{d\theta_{m}} \Big|_{\boldsymbol{\eta}_{i} = \boldsymbol{\eta}_{i}^{*}(\boldsymbol{\theta})} + \frac{d\epsilon_{ij}}{d\eta_{i}} \Big|_{\boldsymbol{\eta}_{i} = \boldsymbol{\eta}_{i}^{*}(\boldsymbol{\theta})} \frac{d\boldsymbol{\eta}_{i}^{*}}{d\theta_{m}} \quad \text{where} \quad \frac{d\epsilon_{ij}}{d\theta_{m}} = \frac{d(\mathbf{d}_{ij} - \hat{\mathbf{y}}_{ij})}{d\theta_{m}} = -\left(\frac{\partial\mathbf{h}}{\partial\theta_{m}} + \frac{\partial\mathbf{h}}{\partial\mathbf{x}_{ij}} \frac{d\mathbf{x}_{ij}}{d\theta_{m}}\right) \\ & + \mathrm{sensitivity differential equations wrt} \, \boldsymbol{\theta}_{m} \end{split}$$

$$\frac{d}{dt} \left(\frac{d\mathbf{x}_i}{d\theta_m} \right) = \frac{\partial \mathbf{f}}{\partial \theta_m} + \frac{\partial \mathbf{f}}{\partial \mathbf{x}_i} \left(\frac{d\mathbf{x}_i}{d\theta_m} \right) \qquad \left(\frac{d\mathbf{x}_i}{d\theta_m} \right) (t_0) = \frac{\partial \mathbf{x}_{0i}}{\partial \theta_m}$$
How to find $\frac{d\eta_i^*}{d\theta}$?
$$\frac{dl_i}{d\eta_i} \Big|_{\eta_i = \eta_i^*(\theta)} = \mathbf{0} \quad \diamondsuit \quad \frac{d}{d\theta} \left(\frac{dl_i}{d\eta_i} \Big|_{\eta_i = \eta_i^*(\theta)} \right) = \mathbf{0} \quad \diamondsuit \quad \frac{d\eta_i^*}{d\theta} = -\left(\frac{d^2l_i}{d\eta_i^2} \Big|_{\eta_i = \eta_i^*(\theta)} \right)^{-1} \frac{d^2l_i}{d\eta_i d\theta} \Big|_{\eta_i = \eta_i^*(\theta)}$$

 $l_i = -\frac{1}{2} \sum_{i=1}^{n_i} \left(\epsilon_{ij}^T \mathbf{R}_{ij}^{-1} \epsilon_{ij} + \log \det(2\pi \mathbf{R}_{ij}) \right) - \frac{1}{2} \eta_i^T \Omega^{-1} \eta_i - \frac{1}{2} \log \det(2\pi \Omega)$

The Inner Optimization Problem

The inner optimization problem consists of finding the η_i that maximizes the individual l_i (for a given θ). Gradient based optimization methods need accurate gradients. The k^{th} component of the gradient of the log-likelihood wrt η_i

$$\frac{dl_i}{d\eta_{ik}} = -\frac{1}{2} \sum_{j=1}^{n_i} \left(2\epsilon_{ij}^T \mathbf{R}_{ij}^{-1} \frac{d\epsilon_{ij}}{d\eta_{ik}} - \epsilon_{ij}^T \mathbf{R}_{ij}^{-1} \frac{d\mathbf{R}_{ij}}{d\eta_{ik}} \mathbf{R}_{ij}^{-1} \epsilon_{ij} + \operatorname{tr} \left[\mathbf{R}_{ij}^{-1} \frac{d\mathbf{R}_{ij}}{d\eta_{ik}} \right] \right) - \eta_i^T \Omega^{-1} \frac{d\eta_i}{d\eta_{ik}}$$

where

 $\frac{d\epsilon_{ij}}{d\eta_{ik}} = \frac{d(\mathbf{d}_{ij} - \hat{\mathbf{y}}_{ij})}{d\eta_{ik}} = -\left(\frac{\partial \mathbf{h}}{\partial \eta_{ik}} + \frac{\partial \mathbf{h}}{\partial \mathbf{x}_{ij}}\frac{d\mathbf{x}_{ij}}{d\eta_{ik}}\right) \quad \text{and} \quad \frac{d\mathbf{R}_{ij}}{d\eta_{ik}} = \frac{\partial \mathbf{R}_{ij}}{\partial \eta_{ik}} + \frac{\partial \mathbf{R}_{ij}}{\partial \mathbf{x}_{ij}}\frac{d\mathbf{x}_{ij}}{d\eta_{ik}}$

The *sensitivity* differential equations wrt η_{ik}

$$\frac{d}{dt} \left(\frac{d\mathbf{x}_i}{d\eta_{ik}} \right) = \frac{\partial \mathbf{f}}{\partial \eta_{ik}} + \frac{\partial \mathbf{f}}{\partial \mathbf{x}_i} \left(\frac{d\mathbf{x}_i}{d\eta_{ik}} \right) \qquad \left(\frac{d\mathbf{x}_i}{d\eta_{ik}} \right) (t_0) = \frac{\partial \mathbf{x}_{0i}}{\partial \eta_{ik}}$$

Second order sensitivities are also required: $\frac{d x_i}{d\eta_{ik}d\theta_m}$ and $\frac{d x_i}{d\eta_{ik}d\eta_{il}}$

Precision, Accuracy, and Performance

Two different levels of magnification of an element of the log-likelihood gradient as a function of the finite difference step, *h*.



Benchmarking – relative estimation times



Model M1: 2-compartment, nonlinear elimination

S-F- η : Sensitivities (inner), Finite differences (outer), improved η starting values

Example: F-F (central diff) to S-S- η gives 50-fold decreased

Starting Values for Random Parameters



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Highlights

- Robust computation of gradients
- Methodology applies to both individual and population log-likelihoods
- Improves computational speed compared to finite differences

References

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